

Likelihood inference for  
a semi-parametric causal model addressing *partial compliance*  
by continuous *principal strata*

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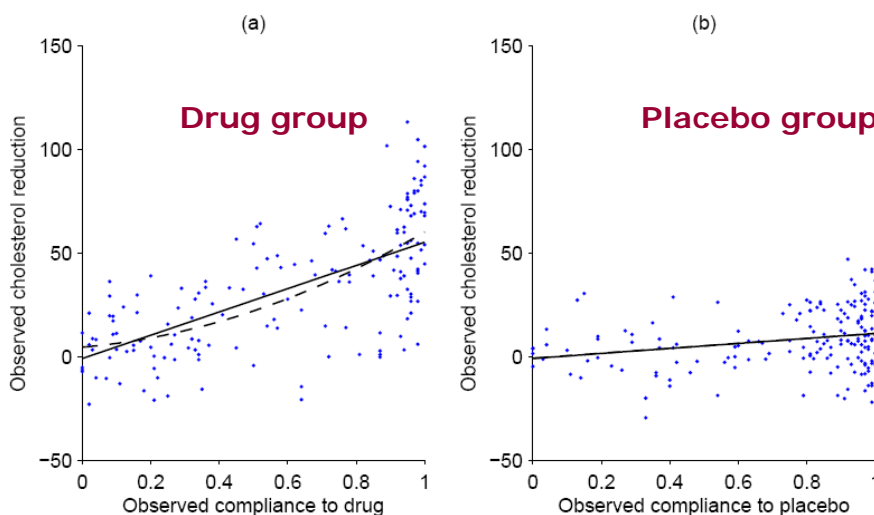
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## 1. Efron-Feldman data

- Subset of data from LRC-CPPT, a *placebo-controlled double-blinded randomized clinical trial* designed to study the effectiveness of *cholestyramine* for lowering cholesterol levels
- data on 335 men: 164 → active pills of the drug 171 → placebo pills
  - binary indicator for treatment assignment
  - proportion of compliance (based on pills taken)
  - continuous outcome variable: average decrease in the cholesterol level during the study (average 7.3 years)



observed compliance to placebo  
larger than  
observed compliance to drug  
(adverse side-effects of the drug)

EF imputed the missing compliances using the percentiles (*equipercentile equating assumption*)

## 2. Modelling strategy

- **EF: EFRON & FELDMAN (1991)**: analysis of a randomized trial with partial non-compliance [Compliance as an Explanatory Variable in Clinical Trials, *JASA* 86, 9-17.]
- **FRANGAKIS & RUBIN (2002)**: principal stratification
  - general framework to deal with non-compliance
  - earliest applications for all-or-none compliance (→ discrete strata)
- **JR: JIN & RUBIN (2008)**: new analysis of Efron & Feldman data using *continuous principal strata* [Principal Stratification for Causal Inference With Extended Partial Compliance, *JASA* 103, 101-111.]
- **BARTOLUCCI & GRILLI (2010)**: new analysis of Efron & Feldman data following the approach of Jin & Rubin but *with a different modeling strategy*
- Treatment indicator  $Z_i$  ( $1 = \text{drug}$ ,  $0 = \text{placebo}$ )
- Potential outcomes (under *Strong access monotonicity*)
  - Compliance:  $d_i$  placebo,  $D_i$  drug
  - Outcome:  $Y_i^{(0)}$  placebo,  $Y_i^{(1)}$  drug
- Principal stratification
  - $(d_i, D_i)$  principal strata (*continuous*)
  - $E(Y_i^{(1)} - Y_i^{(0)} \mid d_i, D_i)$  Principal Causal Effect (PCE)
- Regression models for the outcomes
  - $Y_i^{(0)}$  on  $d_i$  and  $D_i$        $Y_i^{(1)}$  on  $d_i$  and  $D_i$

## 3. Copula for the joint distribution of the compliances

- $d_i$  and  $D_i$  are never jointly observed but some information on their correlation is induced by the equations for the outcomes
- JR used a parametric specification consistent with *Negative side-effect monotonicity*:  $D_i \leq d_i$
- Critical issues:
  - How much information is available on the correlation between the compliances? How sensitive is the inference on the causal effect to the model on the compliances?
- We use a copula for the distribution of  $(d_i, D_i)$  A copula is a flexible way to define a joint distribution from the marginals
- We use a **Plackett copula** which has a single association parameter  $\psi$ 
  - $0 < \psi < 1$       →      negative association
  - $\psi = 1$       →      independence
  - $\psi > 1$       →      positive association
- Advantages of using a copula instead of a parametric density:
  - no constraints on the marginal distributions
  - association captured by a single parameter (to be estimated or used in a sensitivity analysis)

## 4. Model fitting via EM

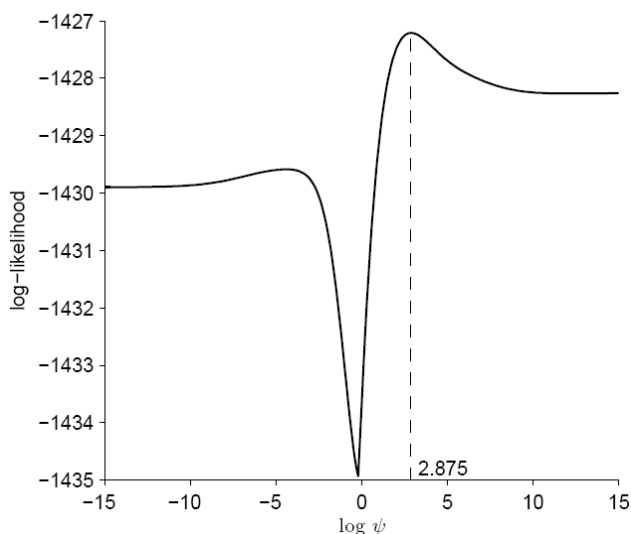
1. Compute the univariate empirical distribution functions of the two compliances  $d_i$  and  $D_i$
2. For a set of values of the association parameter  $\psi$ 
  - i. Estimate the joint distribution function of  $(d_i, D_i)$  using the copula
  - ii. Maximize the likelihood via EM
3. Plot the profile likelihood for  $\psi$

This allows us to see how the different values of  $\psi$  are supported by the data and check for local maxima

- We begin with a general form with quadratic terms, interactions and heteroskedasticity and select via LR test
- FINAL MODEL FOR THE MEANS
- $E(Y_i^{(0)} | d_i, D_i) = -0.269 + 11.243d_i$
- $E(Y_i^{(1)} | d_i, D_i) = -0.269 + 11.243d_i - 21.878D_i + 73.359(d_i \times D_i)$
- **Principal Causal Effect:**  $PCE(d_i, D_i) = (-21.878 + 73.359d_i) D_i$
- The PCE depends on the dose of the drug  $D_i$  and the slope is
  - positive, except when  $d_i < 0.298$  (but this is rare: 12.3% of the subjects in the placebo arm)
  - steeper at higher levels of the placebo compliance  $d_i$

## 5. Maximum likelihood results

Profile log-likelihood for the Plackett association parameter

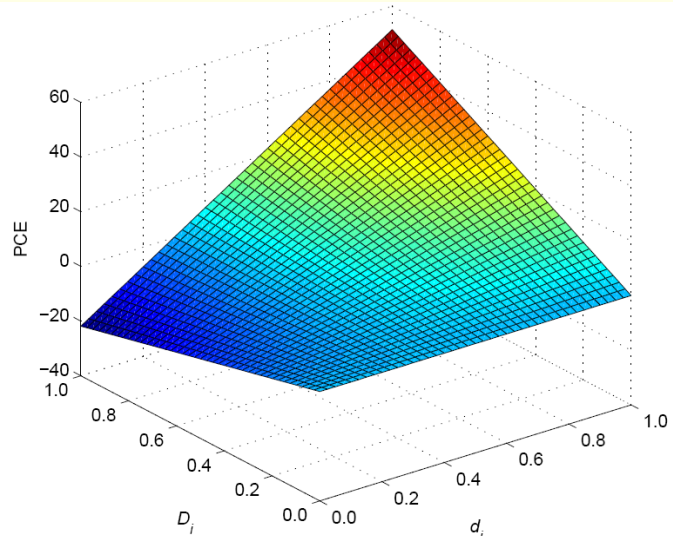


Point estimate of  $\psi$  is 17.727

Independence between  $d_i$  and  $D_i$  (i.e.  $\psi=1$ ) is rejected (p-value < 0.001)

Pearson correlation between  $d_i$  and  $D_i$  is 0.689

Principal Causal Effects (PCE) surface ( $\psi = \text{ML estimate} = 17.727$ )



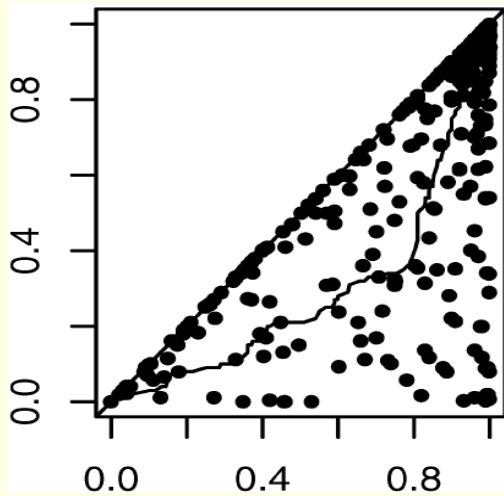
Good agreement with Jin and Rubin

At the median point ( $d_i=0.89$ ,  $D_i=0.70$ ) our ML estimate of PCE is 30.4 with bootstrap CI (22.5, 39.2)

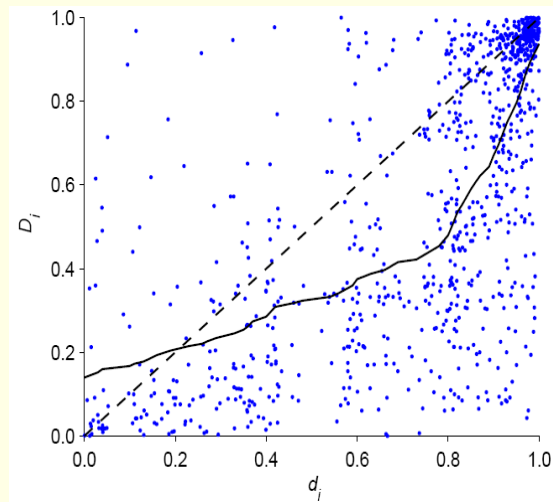
## 6. Association between the compliances

Random draws from the bivariate distribution of the compliances

Jin and Rubin (2008)



Bartolucci and Grilli (2010)



We relax the *negative side-effect monotonicity* (i.e.  $D_i \leq d_i$ )

→ 21.6% of the points go beyond the diagonal, corresponding to individuals with  $D_i > d_i$

## 7. Sensitivity analysis

- The Plackett parameter  $\psi$  determining the association between drug and placebo compliances has *scarce empirical support* and it is identified thanks to the regression equations (which cannot be tested separately)
- The estimate of the PCE depends on the estimate of  $\psi$ : it is not advisable to base the inference exclusively on the point estimate of  $\psi$
- We perform a **sensitivity analysis** to assess how the PCE depends on  $\psi$  (we let it vary in its profile likelihood interval):
  - at the median point ( $d_i=0.89$ ,  $D_i=0.70$ ): PCE  $\in$ (27.4, 34.8)
  - at the Q1-Q3 point ( $d_i=0.59$ ,  $D_i=0.95$ ): PCE  $\in$ (14.0, 29.5)
    - this couple of compliances is unlikely (far from the bulk of the data) → greater sensitivity, in addition to greater sampling variance: the bootstrap CI is very large (-10.8, 34.7)
- Principal Causal Effects are reliably estimated at drug and placebo compliance levels near the sample medians, while inference at unlikely compliance levels appears to be unduly affected by model assumptions